

Lecture #15

Linear Regression

Max. Likelihood :

$$y_i = a_i^T \underline{x} + v_i, \quad i=1, \dots, m$$

measurements known iid $\in \mathbb{R}$ noise.

$$\Leftrightarrow \underline{y}_{m \times 1} = \underline{A}_{m \times n} \underline{x} + \underline{v}_{m \times 1}$$

unknown (deterministic) parameter to be estimated

likelihood f.n. :

$$\prod_{i=1}^m P(y_i - a_i^T \underline{x})$$

$$\Leftrightarrow \text{log-likelihood } l(\underline{x}) = \sum_{i=1}^m \log P(y_i - a_i^T \underline{x})$$

what if $v_i \sim \text{Laplace}(\alpha)$, $\alpha > 0$

$$P_v(z) = \frac{1}{2\alpha} \exp\left(-|z|/\alpha\right)$$

\swarrow Laplace PDF

$$\Leftrightarrow \sum_{i=1}^m \log P(y_i - a_i^T \underline{x}) = - \sum_{i=1}^m \left(2\alpha + \frac{|y_i - a_i^T \underline{x}|}{\alpha} \right)$$

$$\begin{aligned} \therefore \underline{x}_{ML}^* &= \operatorname{argmax} \ell(\underline{x}) \\ &= \operatorname{argmin}_{\underline{x} \in \mathbb{R}^n} \|\underline{y} - A\underline{x}\|_1 \quad (\text{LP}) \end{aligned}$$

Many PDFs in practice are log-concave

e.g. multivariate Gaussian, Laplace, Exponential, Uniform over convex set, Beta, Gamma, χ^2 etc.

log-concavity is preserved under

① product

② Marginalization

③ Convolution

$$\int p(x, y) dx = q(y)$$

$$\int p(x, y) dy = r(x)$$

Conversely, we can give statistical interpretation of any convex optimization problem of the form:

$$\underline{x}^* = \underset{\underline{x} \in \mathbb{R}^n}{\operatorname{argmin}} \sum_{i=1}^m \phi(b_i - a_i^T \underline{x})$$

This is an ML problem where

$v_i \rightarrow$ iid. with PDF $p_v(z) \propto \exp(-\phi(z))$

$$p_v(z) = \frac{\exp(-\phi(z))}{\int \exp(-\phi(z)) dz}$$

ML with Poisson's distribution:

$Y \in \mathbb{Z}_+ = \{0, 1, 2, \dots\}$ (Poisson distribution)
↑ random variable
→ distribution Poisson(μ)

$$P(Y = k) = \frac{\exp(-\mu) \mu^k}{k!}$$

$$\mu = a^T u + b$$

↑ parameters

↑ explanatory variable

Given (u_i, y_i)

↑ explanatory variable measured.

↑ observed/realized measurement

log-likelihood : $l(\underline{a}, b)$

$$= \sum_{i=1}^m \left\{ y_i \log(\underline{a}^T \underline{u}_i + b) - (\underline{a}^T \underline{u}_i + b) - \log(y_i!) \right\}$$

$$\underbrace{(\underline{a}_{ML}^*, b_{ML}^*)}_{\text{Max. likelihood estimate of the parameters } \underline{a} \text{ \& } b} = \underset{\substack{\underline{a} \in \mathbb{R}^n \\ b \in \mathbb{R}}}{\text{argmax}} \underbrace{\left\{ y_i \log(\underline{a}^T \underline{u}_i + b) - (\underline{a}^T \underline{u}_i + b) \right\}}_{\text{convex fcn}}$$

ML estimate of covariance of some Gaussian random vector.

$$\underline{y} \in \mathbb{R}^n, \quad \underline{y} \sim \mathcal{N}(\underline{0}, R), \quad R \in \mathbb{S}_{++}^n$$

$$\therefore P(\underline{y}) = \frac{1}{\sqrt{(2\pi)^n \det(R)}} \exp\left(-\frac{1}{2} \underline{y}^T R^{-1} \underline{y}\right)$$

We want to estimate the matrix $R \succ 0$ based on N samples $\underline{y}_1, \dots, \underline{y}_N \in \mathbb{R}^n$ drawn from this distribution.

Log-likelihood:

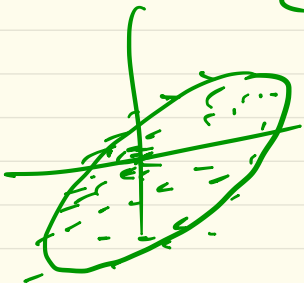
$$l(R) = \log P_R(\underline{y}_1, \dots, \underline{y}_N)$$

$$\Rightarrow \ell(R) = - \binom{Nn}{2} \log(2\pi) - \binom{N}{2} \log \det(R) - \frac{1}{2} \sum_{k=1}^N \underline{y}_k^T R^{-1} \underline{y}_k$$

where

$$Y := \frac{1}{N} \sum_{k=1}^N \underline{y}_k \underline{y}_k^T \quad (\text{sample covariance matrix})$$

$$\text{covar}(X) = \mathbb{E} \left(\begin{array}{c} (X - \mathbb{E}(X)) \\ (X - \mathbb{E}(X))^T \end{array} \right)$$



Can show:

$\ell(R)$ is concave in R when $R \preceq 2Y$ (only on a subset of S_{++}^n)

change-of-variable:

$$S := R^{-1} \succ 0$$

Information matrix \quad cov. matrix

(Recall that matrix inversion preserves pos. definiteness)

$$\begin{aligned} \therefore \ell(S) = & -\left(\frac{Nn}{2}\right) \log(2\pi) + \underbrace{\left(\frac{N}{2}\right) \log \det(S)}_{\text{concave}} \\ & - \underbrace{\left(\frac{N}{2}\right) \text{tr}(SY)}_{\text{concave}} \end{aligned}$$

Concave in S

$$\begin{aligned} \therefore S_{ML}^* &= \underset{S \in \mathbb{S}_{++}^n}{\text{argmax}} \log \det(S) - \text{tr}(SY) \\ &= \underset{S \succ 0}{\text{argmin}} -\log \det(S) + \text{tr}(SY) \end{aligned}$$

convex optimization problem

∴ Analytical soln:

$$S_{ML}^* = Y^{-1}$$

ML estimate
of ensemble
covariance

sample
covariance

$$\Rightarrow (R_{ML}^*)^{-1} = Y^{-1} \Leftrightarrow R_{ML}^* = Y$$

We can handle additional constraints on R or S

e.g. $L \preceq R \preceq U \Leftrightarrow U^{-1} \preceq R^{-1} \preceq L^{-1}$

$\Rightarrow \begin{cases} \max \ell(S) \\ S \succeq 0 \\ \& U \preceq S \preceq L^{-1} \end{cases}$ ← LMI

(operator
monotonicity,
HW 3, p 2(e))

e.g.

$$\frac{\lambda_{\max}(R)}{\lambda_{\min}(R)} \leq \kappa_{\max} \left. \begin{array}{l} \text{condition number} \\ \text{constraint on} \\ \text{matrix } R \end{array} \right\}$$

$$\Leftrightarrow \exists u > 0 \text{ s.t. } uI \preceq S \preceq \kappa_{\max} uI$$

$$\therefore S_{ML}^* = \operatorname{argmax}_S l(S)$$

$$S \succ 0, u \in \mathbb{R}$$

$$\text{s.t. } uI \preceq S \preceq k_{\max} uI$$

} SDP

MAP (Maximum a posteriori Probability Estimation)

x ← to be estimated } Joint PDF

y ← observed variable } $l(\underline{x}, \underline{y})$

(In ML estimation:

x is deterministic, not a random variable)

Prior of x : $P_X(x) = \int P(x, y) d\underline{y}$

(before we observe y)

$$P_Y(y) = \int P(x, y) d\underline{x}$$

$$P_{Y|X}(\underline{x}, \underline{y}) = \frac{P(\underline{x}, \underline{y})}{P_X(\underline{x})} \quad \left. \vphantom{\frac{P(\underline{x}, \underline{y})}{P_X(\underline{x})}} \right\} \text{Bayes Rule}$$

$$\Rightarrow P(\underline{x}, \underline{y}) = \boxed{P_{Y|X}(\underline{x}, \underline{y}) P_X(\underline{x})}$$

$$P_{X|Y}(\underline{x}, \underline{y}) = \frac{P(\underline{x}, \underline{y})}{P_Y(\underline{y})} = \frac{P_{Y|X}(\underline{x}, \underline{y}) P_X(\underline{x})}{P_Y(\underline{y})}$$

$$\begin{aligned} \underline{x}_{\text{MAP}}^* &= \underset{\underline{x} \in \mathbb{R}^n}{\text{argmax}} P_{X|Y}(\underline{x}, \underline{y}) \\ &= \underset{\underline{x} \in \mathbb{R}^n}{\text{argmax}} \frac{P_{Y|X}(\underline{x}, \underline{y}) P_X(\underline{x})}{P_Y(\underline{y})} \\ &= \underset{\underline{x} \in \mathbb{R}^n}{\text{argmax}} \underbrace{P(\underline{x}, \underline{y})}_{\text{joint}} \end{aligned}$$

Taking $\log(\cdot)$

$$\hat{\underline{x}}_{\text{MAP}} = \underset{\underline{x} \in \mathbb{R}^n}{\text{argmax}} \left\{ \underbrace{\log p_{Y|X}(\underline{x}, \underline{y})}_{\substack{\text{log-likelihood} \\ \text{of conditional}}} + \underbrace{\log p_X(\underline{x})}_{\substack{\text{penalizing} \\ \text{unlikely} \\ \underline{x}}} \right\}$$

e.g. linear regression (with MAP)
with iid. noise

$$y_i = \underline{a}_i^T \underline{x} + v_i, \quad i=1, \dots, m$$

(iid with PDF p_v on \mathbb{R})

This time

$$\underline{x} \sim p_X(\underline{x}) \text{ on } \mathbb{R}^n \text{ (Prior PDF)}$$

$$\therefore \text{Joint } p(\underline{x}, \underline{y}) = p_X(\underline{x}) \prod_{i=1}^m p_v(y_i - \underline{a}_i^T \underline{x})$$

$$\therefore \hat{\underline{x}}_{\text{MAP}}^* = \underset{\underline{x} \in \mathbb{R}^n}{\text{argmax}} \left\{ \log p_x(\underline{x}) + \sum_{i=1}^m \log p_v(y_i - a_i^T \underline{x}) \right\}$$

extra term
compared to ML

e.g.

If $v_i \sim \text{Uniform}([-a, a])$

& $\underline{x} \sim \mathcal{N}(\underline{\mu}, \Sigma)$

then

this reduces to

$$\hat{\underline{x}}_{\text{MAP}}^* = \underset{\underline{x} \in \mathbb{R}^n}{\text{argmin}} \left(\underline{x} - \underline{\mu} \right)^T \Sigma^{-1} \left(\underline{x} - \underline{\mu} \right) \left. \vphantom{\argmin} \right\} \text{QP}$$

s.t. $\| A \underline{x} - \underline{y} \|_{\infty} \leq a$

Non-parametric Estimation of Prob. Distribution over a finite set $\{\alpha_1, \dots, \alpha_n\} \subset \mathbb{R}$



Probability (discrete) distribution $\underline{p} \in \mathcal{S}$
Probability simplex

$$\mathcal{S} = \left\{ \underline{p} \in \mathbb{R}^n_{\geq 0} \mid \mathbf{1}^T \underline{p} = 1 \right\}$$

Notice that.

given $f: \mathbb{R} \rightarrow \mathbb{R}$,

(any nonlinear function f)

$$\mathbb{E}[f(x)] = \sum_{i=1}^n p_i f(\alpha_i)$$

linear in \underline{p} .

e.g. $P(X \in \mathcal{C}) = \underbrace{c^T p}_{\text{linear in } p}$ where $c_i = \begin{cases} 1 & \text{if } x_i \in \mathcal{C} \\ 0 & \text{otherwise} \end{cases}$

e.g. $E[X] = a$ (given) $\quad E[X^2] = b$ (given)

$$\sum_{i=1}^n \alpha_i p_i = a$$

$$\Leftrightarrow \sum_{i=1}^n \alpha_i^2 p_i = b$$

$$P(X \geq 0) \leq 0.3$$

$$\Leftrightarrow \sum_{\alpha_i \geq 0} p_i \leq 0.3$$

Linear in p

e.g. Nonlinear in vector \underline{p}

$$\text{Var}(X) = c$$

$$\begin{aligned} & \parallel \\ & \mathbb{E}[X^2] - (\mathbb{E}X)^2 \\ & = \sum_{i=1}^n \alpha_i^2 p_i - \left(\sum_{i=1}^n \alpha_i p_i \right)^2 \end{aligned}$$

quadratic in \underline{p}

$$\text{Var}(X) := \mathbb{E}[(X - \mathbb{E}(X))^2]$$

$$= \mathbb{E}[X^2 - 2X(\mathbb{E}X) + (\mathbb{E}X)^2]$$

$$= \mathbb{E}[X^2] - [2(\mathbb{E}X)^2 + (\mathbb{E}X)^2]$$

Max. Entropy Problem

$$\max_{\underline{p} \in \mathcal{S}} \text{Entropy} \Rightarrow \max_{\underline{p} \in \mathcal{S}} - \sum_{i=1}^n p_i \log p_i$$

↖ concave

$$\min_{\underline{p} \in \mathcal{S}} \sum_{i=1}^n p_i \log p_i$$

↕

convex.

s.t. $\underline{p} \in \mathcal{C}$ ← some convex set

Min. KL Divergence Problem: Kullback-Leibler (KL) divergence

$$\min_{\underline{p} \in \mathcal{S}} \left[\sum_{i=1}^n p_i \log \left(\frac{p_i}{q_i} \right) \right] = D_{KL}(\underline{p} \parallel \underline{q})$$

s.t. $\underline{p} \in \mathcal{C}$
↕
convex set

If $\underline{q} = \left(\frac{1}{n} \mathbb{1} \right)$ (uniform distribution)
then $D_{KL}(\underline{p} \parallel \underline{q}) = \sum_{i=1}^n p_i \log p_i$